

Machine learning theory for time series

Exponential inequalities for nonstationary Markov chains

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CIMFAV seminar
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- 1 Short introduction to machine learning theory
- 2 Machine learning and time series
 - Machine learning & stationary time series
 - Nonstationary Markov chains

Generic machine learning problem

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Sub-gamma random variables

Definition

T is said to be sub-gamma iff
 $\exists(v, w)$ such that $\forall k \geq 2$,

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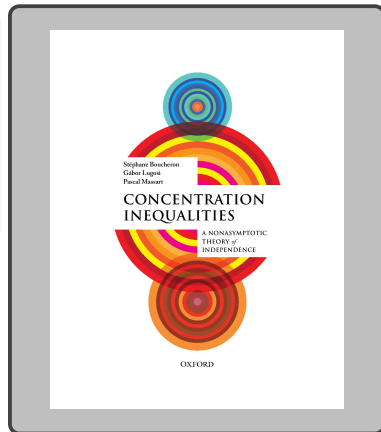
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- any Z with
 $\mathbb{P}(|Z| \geq t) \leq \mathbb{P}(|T| \geq t)$.



Bernstein's inequality

Theorem

Let $T_1, \dots, T_n$ be i.i.d and (v, w) -sub-gamma random variables. Then, $\forall \zeta \in (0, 1/w)$,

$$\mathbb{E} \exp \left(\zeta \sum_{i=1}^n [T_i - \mathbb{E} T_i] \right) \leq \exp \left(\frac{nv\zeta^2}{2(1 - w\zeta)} \right).$$

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$$\mathbb{P} \left[|R(\theta) - r(\theta)| > t \right] \leq 2 \exp \left[\frac{vs^2}{2(n - ws)} - st \right]$$

Finite set of predictors and union bound

$$\mathbb{P}\left[\exists \theta \in \Theta, |R(\theta) - r(\theta)| > t\right] = \mathbb{P}\left[\bigcup_{\theta \in \Theta} \{|R(\theta) - r(\theta)| \geq t\}\right]$$

Finite set of predictors and union bound

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On the complement,

$$R(\hat{\theta}) \leq \min_{\theta \in \Theta} R(\theta) + \frac{vs}{n - ws} + \frac{2 \log \frac{2|\Theta|}{\alpha}}{s}$$

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On the complement, for
 $s = \lceil n/(2w) \rceil \wedge \sqrt{(n/v) \log(2|\Theta|/\alpha)},$

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With probability at least $1 - \alpha$,

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With proba. at least $1 - \alpha$, with α not ridiculously small,

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Infinite parameter set

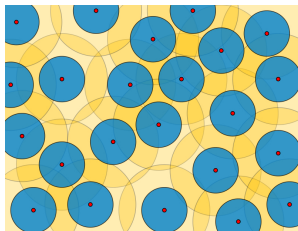
Θ compact $\Rightarrow \exists$ a finite $\Theta^{(\varepsilon)}$ such that

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Infinite parameter set (image from Wikipedia)

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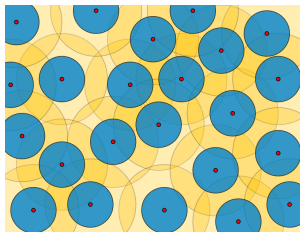
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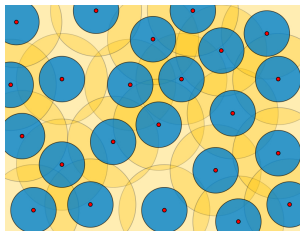
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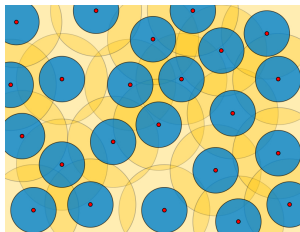
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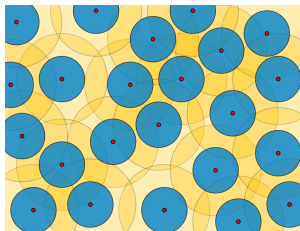
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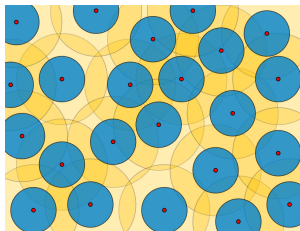
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With proba. at least $1 - \alpha$,

$$R(\hat{\theta}) \leq \min_{1 \leq m \leq M} \left\{ \min_{\theta \in \Theta_m} R(\theta) + \sqrt{\frac{d_m}{n}} \left(2L + 2\sqrt{v \log \frac{2nM}{d\alpha}} \right) \right\}.$$

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- removing $\log(n)$ by refinement of the ε -net structure.
- faster rates : $\sqrt{d/n}$ becomes d/n thanks to a better analysis of v under the **Bernstein condition**.
- relaxing the sub-gamma assumption.
- more flexible way to measure the complexity of Θ : **PAC-Bayesian bounds**.

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An example on Markov chains

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Stochastic Processes and their Applications

Volume 125, Issue 1, January 2015, Pages 60-90



Deviation inequalities for separately Lipschitz functionals of iterated random functions

Jérôme Dedecker ^a  , Xiequan Fan ^b 

An example on Markov chains



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Assume, for $\rho \in [0, 1)$ and $C > 0$,

$$\begin{aligned} \mathbb{E}[d(F(x, \varepsilon_1), F(x', \varepsilon_1))] &\leq \rho d(x, x') \\ d(F(x, y), F(x, y')) &\leq C\delta(y, y'). \end{aligned}$$

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Define

$$\hat{\theta} = \operatorname{argmin}_{\theta \in \Theta} r(\theta).$$

Notations in DF15

Define $G_{X_1}(x) = \int d(x, x')P_{X_1}(dx'),$
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$$\text{Define } \mathcal{V} = \frac{V_1 + V_2}{1 - \rho^2}, \quad \delta = \frac{1 - \rho}{M}.$$

Dedecker and Fan's inequality

Theorem (Dedecker & Fan 2015)

Consider a separately Lipschitz function $f : \mathcal{X}^n \rightarrow \mathbb{R}$:

$$|f(x_1, \dots, x_n) - f(x'_1, \dots, x'_n)| \leq \sum_{t=1}^n d(x_t, x'_t).$$

Then, for any $s \in [0, \delta^{-1})$,

$$\mathbb{E} \left[e^{\pm s \{f(X_1, \dots, X_n) - \mathbb{E}[f(X_1, \dots, X_n)]\}} \right] \leq \exp \left(\frac{(n-1)s^2 \mathcal{V}}{2(1-s\delta)} \right).$$

Consequences of DF15 for prediction

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$$\text{Take } f(X_1, \dots, X_n) = \frac{1}{L} \sum_{i=2}^n \ell(f_{\theta}(X_{i-1}) - X_i).$$

Then for any $0 \leq s < (n-1)/(L(1+\rho)\delta)$,

$$\begin{aligned} \mathbb{P} \left[|R(\theta) - r(\theta)| > t \right] \\ \leq 2 \exp \left(\frac{s^2(1+\rho)^2 L^2 \mathcal{V}}{2(n-1) - 2s(1+\rho)\delta L} - st \right). \end{aligned}$$

Learning theorem for Markov chains

Assume $|\Theta^{(\varepsilon)}| \leq \varepsilon^{-d}$.

Theorem

As soon as $n \geq 1 + 4\delta^2 d \log(Ln)/\mathcal{V}$ we have, with probability at least $1 - \alpha$,

$$R(\hat{\theta}) \leq \inf_{\theta \in \Theta} R(\theta) + C_1 \sqrt{\frac{d \log(Ln)}{n-1}} + C_2 \frac{\log\left(\frac{4}{\alpha}\right)}{\sqrt{n-1}} + \frac{C_3}{n},$$

where $C_1 = 4(1 + \rho)L\sqrt{\mathcal{V}}$, $C_2 = 2(1 + \rho)L\sqrt{\mathcal{V}} + 2\delta$ and $C_3 = 3[G_{\varepsilon}(0) + G_{X_1}(0)]/(1 - \rho) + \mathcal{V}/(2\delta)$.

Other works on ML & TS

Study of $X_t = F(\varepsilon_t; X_{t-1}, X_{t-1}, \dots)$ in



P. Alquier, O. Wintengerger. Model selection for weakly dependent time series forecasting. *Bernoulli*, 2012.

based on Rio's version of Hoeffding's inequality



E. Rio. Inégalités de Hoeffding pour les fonctions lipschitziennes de suites dépendantes. *CRAS*, 2000.

Rates in $\sqrt{\frac{d}{n}}$.

Fast rates

Rates in $\frac{d}{n}$ for quadratic ℓ in



P. Alquier, X. Li, O. Wintengerger. Prediction of time series by statistical learning : general losses and fast rates.. *Dependence Modeling*, 2013.

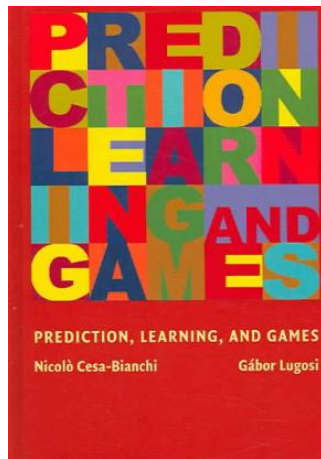
based on Samson's version of Bernstein's inequality for φ -mixing processes



P.-M. Samson. Concentration of measure inequalities for markov chains and φ -mixing processes. *The Annals of Probability*, 2000.

Online prediction approach

The online prediction approach provides tools to aggregate predictors without stochastic assumptions on the data.



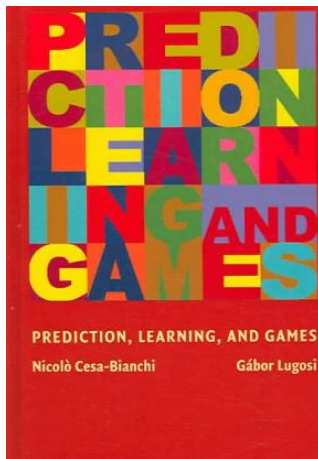
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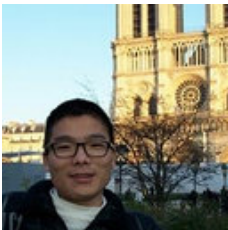
C. Giraud, F. Roueff, A. Sanchez-Perez. Aggregation of predictors for nonstationary sub-linear processes and online adaptive forecasting of time varying autoregressive processes. *The Annals of Statistics*, 2015.

takes advantage of this approach to predict time varying AR processes.



- 1 Short introduction to machine learning theory
- 2 Machine learning and time series
 - Machine learning & stationary time series
 - Nonstationary Markov chains

Nonstationary Markov chains



P. Alquier, P. Doukhan, X. Fan.
Exponential inequalities for
nonstationary Markov Chains.
Preprint arxiv :1808.08811, 2018.



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Example 2 : T -periodic AR(1)

$$X_t = a_{t[T]} X_{t-1} + \varepsilon_t, \quad \max_{1 \leq t \leq T} |a_t| \leq \rho.$$

Example : T -periodic AR(1)

4-periodic AR(1), $(a_1, a_2, a_3, a_4) = (0.8, 0.5, 0.9, -0.7)$.

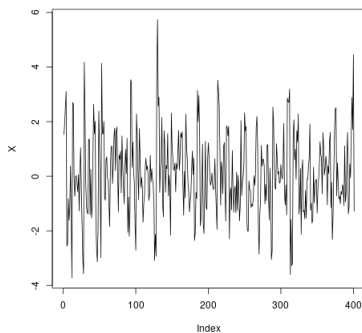


Figure – Simulated data.

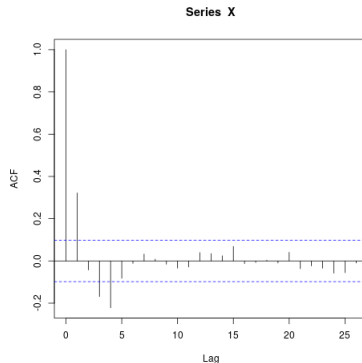


Figure – Autocorrelations.

Bernstein's inequality

Theorem (ADF18)

Assume that, for any $k \geq 2$,

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Consider a separately Lipschitz function $f : \mathcal{X}^n \rightarrow \mathbb{R}$. For any $s \in [0, \delta^{-1})$,

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Problem : estimation of the (best) period

From now, assume that

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Let $(f_\theta, \theta \in \Theta)$ be a set of predictors $\mathcal{X} \rightarrow \mathcal{X}$, define $\mathcal{H}(\varepsilon)$ as $\log |\Theta^{(\varepsilon)}|$. Put, for any T and $\theta_{1:T} = (\theta_1, \dots, \theta_T) \in \Theta^T$:

$$r_n(\theta_{1:T}) = \frac{1}{n-1} \sum_{i=2}^n \ell(f_{\theta_{i[T]}}(X_{i-1}) - X_i)$$

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if actually $f_t^* = f_{t[T]}^*$, where $C_0 = L(1 + \rho) \left[\frac{G_\varepsilon(0)}{1-\rho} + G_{X_1}(0) \right]$.

Estimators

Estimation for a given period T :

$$\hat{\theta}_{1:T} = (\hat{\theta}_1, \dots, \theta_T) = \underset{\theta_{1:T}=(\theta_1, \dots, \theta_T)}{\operatorname{argmin}} r_n(\theta_{1:T}).$$

Period selection :

$$\hat{T} = \underset{1 \leq T \leq T_{\max}}{\operatorname{argmin}} \left[r_n(\hat{f}_{1:T}) + \frac{C_1}{2} \sqrt{\frac{T \mathcal{H}(\frac{1}{Ln})}{n-1}} \right]$$

where C_1 is as in the stationary case : $C_1 = 4(1 + \rho)L\sqrt{\mathcal{V}}$.

Analysis of the estimators

Theorem (ADF18)

As soon as $n \geq 1 + 4\delta^2 T_{\max} \mathcal{H}(\frac{1}{Ln})/\mathcal{V}$, with probability at least $1 - \alpha$,

$$R(\hat{\theta}_{1:\hat{T}}) \leq \inf_{1 \leq T \leq T_{\max}} \inf_{\theta_{1:T} \in \Theta^T} \left[R(\theta_{1:T}) \right. \\ \left. + C_1 \sqrt{\frac{T \mathcal{H}(\frac{1}{Ln})}{n-1}} + C_2 \frac{\log\left(\frac{4T_{\max}}{\alpha}\right)}{\sqrt{n-1}} + \frac{C_3}{n} \right].$$

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In practice, C_1 , C_2 and C_3 are too large and ρ is not known anyway... we recommend to use the slope heuristic here.

Slope heuristic

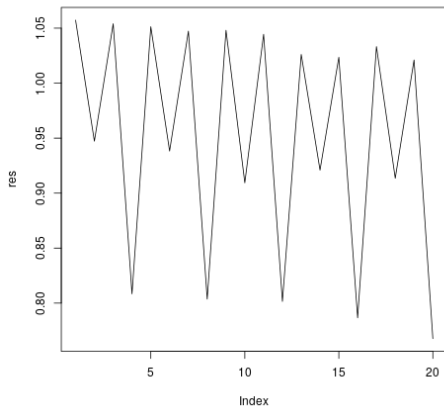


Figure – Empirical risk as a function of T .

Thank you !